

Total Points: 26

1. For each of the following random variables, identify the distribution that most closely matches the situation. Justify your choice and describe any assumptions that you made.

a. (4pts) On average, a flight from Houston to Dallas leaves every 60 minutes. Let X_1 be the amount of time I have to wait for the next airplane to come

ANSWER: $X_1 \sim \text{exponential}(60)$ (1pt) for exponential (1pt) for $\theta = 60$

Reason: X_1 is the waiting time for a poisson r.v. (plane arrivals) with an average waiting time of 60. (1pt) for a reason

Assumptions: Plane arrival times are independent of each other, etc. (1pt) for an assumption

b. (4pts) 12oz. beer bottles have a mean volume of 11.8oz and a variance of 0.7oz. Let X_2 be the amount of beer in the bottle from this sample.

ANSWER: $X_2 \sim \text{normal}(11.8, 0.7)$ or truncated normal (1pt) for normal (1pt) for $\mu = 11.8$ $\sigma^2 = 0.7$ ($\sigma = 0.837$)

Reason X_2 is symmetrically distributed around 11.8oz (we'd assume) and has a higher chance of being close to 11.8oz than far from it. (1pt) a reason

Assumptions - we must assume that a bottle could hold both $-\infty$ and ∞ oz. Or we must truncate our distribution to reasonable values. (1pt) on assumption

C. (4pts) The registrar gets lazy and decides to use a random number generator to determine the threshold GPA for the presidents honor roll. Let X_3 be the minimum GPA to get you on the honor roll.

ANSWER

$$X_3 \sim \text{Uniform}(0.0, 4.0)$$

(1pt) uniform.

(1pt) $a=0$
 $b=4$

reason: each possible gpa has an equal chance of being selected. GPAs fall between 0 and 4.

(1pt) a reason

Assumption: the random number generator uses the same range as GPAs, it doesn't weight the possibility of choosing values, etc.

(1pt) an assumption

2. For each of the following random variables, find the specified probability using the CDF.

a. (2pts) $X_1 \sim \text{exponential}(\theta=10)$ $P(10 < X_1 < 100)$

ANSWER: $P(10 < X_1 < 100) = F(100) - F(10)$

(1pt) translating probability to integral of CDF

$$= 1 - e^{-\frac{100}{10}} - (1 - e^{-\frac{10}{10}})$$

(1pt) correctly supplying CDF with $\theta=10$

$$= 1 - e^{-10} - 1 + e^{-1}$$

$$= e^{-1} - e^{-10} \approx 0.368$$

2b (2pts) $X_2 \sim \text{Gamma}(\alpha=1, \beta=2), P(1 < X_2 < 5)$

ANSWER: The CDF of the gamma distribution does not have a closed form. But that does not mean it cannot be calculated. For example, with Wolfram Alpha:

$$\begin{aligned}
 P(1 < X_2 < 5) &= F(5) - F(1) && \text{1pt} \\
 &= \text{CDF}[\text{gamma distribution}[1, \frac{1}{2}, 5]] - \text{CDF}[\text{gamma} \dots] && \text{1pt} \\
 &= \frac{1}{e^2} - \frac{1}{e^{10}} \approx 0.135 && \text{OR}
 \end{aligned}$$

Alternatively, students could integrate the pdf. We will reward full credit for doing this here, in part b, only.

$$\begin{aligned}
 P(1 < X_2 < 5) &= \int_1^5 \frac{2^1}{\Gamma(1)} x^{1-1} e^{-2x} dx && \text{+1pt} \\
 &= \int_1^5 2e^{-2x} dx && \text{+1pt} \\
 &= \frac{e^{-2} - e^{-10}}{e^{-2}} \approx 0.135 && \text{OR}
 \end{aligned}$$

Or Bonus students may recognize that $\text{Gamma}(\alpha=1, \beta=2)$ is the same as exponential ($\theta = \frac{1}{2}$) and use the exponential CDF

$$\begin{aligned}
 P(1 < X_2 < 5) &= F(5) - F(1) && \text{+1pt for recognizing} \\
 &= 1 - e^{-2 \cdot 5} - (1 - e^{-2 \cdot 1}) && \text{+1pt} \\
 &= e^{-2} - e^{-10} \approx 0.135 && \text{+1pt}
 \end{aligned}$$

c. (2pts) $X_3 \sim \text{Normal}(\mu=0, \sigma^2=10), P(-10 < X_3 < 10)$

ANSWER: The cdf of a normal distribution is not easily computed. However, we can look up values of a standard normal distribution in a probability table. To do that here, we should first convert X_3 to a standard normal r.v. ($\mu=0, \sigma^2=1$)

$$\begin{aligned}
 P(-10 < X_3 < 10) &= P\left(-\frac{10-\mu}{\sigma} < \frac{X_3-\mu}{\sigma} < \frac{10-\mu}{\sigma}\right) = P\left(-\frac{10}{\sqrt{10}} < Z < \frac{10}{\sqrt{10}}\right) && \text{1pt for converting to Z} \\
 &= P(-\sqrt{10} < Z < \sqrt{10}) \\
 &= \Phi(\sqrt{10}) - \Phi(-\sqrt{10}) && \text{1pt for correctly interpreting probability} \\
 &\approx 0.999 - 0.000782 \\
 &\approx 0.998
 \end{aligned}$$

3. Given that $X \sim \text{Exponential}(\theta)$, find the pdfs of the following two transformations of X . Do they correspond to named distributions that we know about?

a. (4pts) $Y = X^2$

notice that the range of the exponential distribution is $X \geq 0$.

Answer: $\therefore u(x) = x^2$ is invertible with inverse, $v(x) = \sqrt{x}$

(1pt) for checking invertibility

Hence, we can use the change of variables formula

$$f_Y(y) = f_X(v(y)) |v'(y)|$$

$$= \frac{1}{\theta} e^{-\frac{v(y)}{\theta}} |v'(y)|$$

$$= \frac{1}{\theta} e^{-\frac{\sqrt{y}}{\theta}} |v'(y)|$$

$$= \frac{1}{\theta} e^{-\frac{\sqrt{y}}{\theta}} \frac{1}{2\sqrt{y}}, y \in [0, \infty)$$

(1pt) for correctly defining $f_Y(y)$ in terms of $f_X(x)$

(2pt) for correctly identifying $v(y)$ and $v'(y)$

(1pt) for correctly specifying range

Does not correspond to a known distribution

(1/2pt) for recognizing not a dist

ALTERNATIVE ANSWER: Or, we can use the distribution function technique

1. Figure out the range

$$x \in [0, \infty) \Rightarrow y = x^2 \in [0, \infty)$$

2. Write the CDF for Y , then substitute for x

$$F_Y(y) = P(Y < y)$$

$$= P(x^2 < y)$$

3. Find the region in X such that $u(x) < y$

$$P(x^2 < y) = P(-\sqrt{y} < x < \sqrt{y})$$

or

$$P(x < \sqrt{y}) \text{ because } y \in [0, \infty)$$

(1pt) for correctly specifying $F_Y(y)$ in terms of x

4. Integrate to compute CDF

$$P(-\sqrt{y} < x < \sqrt{y}) = \int_{-\sqrt{y}}^{\sqrt{y}} f(x) dx \quad \left(\text{or } \int_0^{\sqrt{y}} f(x) dx \right)$$

$$= \int_{-\sqrt{y}}^0 f(x) dx + \int_0^{\sqrt{y}} f(x) dx$$

$$= 0 + \int_0^{\sqrt{y}} f(x) dx \quad \text{because } f(x) = 0 \quad \forall x < 0 \quad \text{for an exponential dist.}$$

$$= \int_0^{\sqrt{y}} \frac{1}{\theta} e^{-\frac{x}{\theta}} dx$$

$$= 1 - e^{-\frac{\sqrt{y}}{\theta}} \quad y \in [0, \infty)$$

(1pt) for correctly transforming probability into an integral of $f(x)$

5. Differentiate to get the pdf

$$\frac{d}{dy} \left(1 - e^{-\frac{\sqrt{y}}{\theta}} \right) = \frac{1}{2\theta\sqrt{y}} e^{-\frac{\sqrt{y}}{\theta}} \quad y \in [0, \infty)$$

(1/2 pt) for finding the pdf by differentiating

(1pt) for correctly specifying the range of y

This does not correspond to a known distribution

(1/2) for recognizing this

b. (4pts) $Z = e^x$

ANSWER 1: With the distribution function method (the steps are as listed above).

1.) $x \in [0, \infty) \quad y = e^x \in [1, \infty)$

2.) $F_Y(y) = P(y < Y) = P(e^x < y)$

3.) $= P(x < \ln(y))$

4.) $= \int_0^{\ln(y)} f(x) dx$

$$= \int_0^{\ln(y)} \frac{1}{\theta} e^{-\frac{x}{\theta}} dx$$

$$= 1 - e^{-\frac{\ln(y)}{\theta}}$$

(1pt) for correctly specifying $F_Y(y)$ in terms of x

(1pt) for correctly transforming probability into an integral

$$5. \frac{d}{dy} \left[1 - e^{-\frac{\ln y}{\theta}} \right] = \frac{1}{\theta} y^{-\left(\frac{\theta+1}{\theta}\right)}$$

$\frac{1}{2}$ pt for finding the pdf by differentiation

$$= \frac{1}{\theta} y^{-\frac{1}{\theta}-1}$$

$y \in [1, \infty)$ 1 pt for specifying the new range

This is not a distribution that we know

$\frac{1}{2}$ pt for recognizing not a known dist

Alternatively, students could use the change of variables method

as $u(x) = e^x$ is invertible with $v(x) = \ln(x)$

1 pt for checking invertibility

$$f_y(y) = f_x(v(y)) |v'(y)|$$

$$= \frac{1}{\theta} e^{-\frac{v(y)}{\theta}} |v'(y)|$$

$$= \frac{1}{\theta} e^{-\frac{\ln(y)}{\theta}} |v'(y)|$$

$$= \frac{1}{\theta} e^{-\frac{\ln(y)}{\theta}} \left| \frac{1}{y} \right|$$

$$= \frac{1}{\theta} y^{-\frac{1}{\theta}-1}$$

$$= \frac{1}{\theta} y^{-\frac{1}{\theta}-1}$$

$$y \in [1, \infty)$$

1 pt for correctly defining $f_y(y)$

$\frac{1}{2}$ pt for correctly identifying $v(y)$ and $v'(y)$

$\frac{1}{2}$ pt for recognizing not a known dist

1 pt for correctly specifying range

4 (Bonus 3pts) Show that if $X \sim \text{Normal}(\mu, \sigma^2)$ then $Z = \frac{(X - \mu)}{\sigma}$ has a standard normal distribution.

ANSWER: Recall from lecture, if $X \sim \text{Normal}(\mu, \sigma^2)$ and $Y = aX + b$, THEN

$$Y \sim \text{Normal}(a\mu + b, a^2\sigma^2)$$

(**)

$$\text{So, } Z = \frac{(X - \mu)}{\sigma} = \frac{1}{\sigma} X - \frac{\mu}{\sigma}$$

2 pts for solution shown

$$\sim \text{Normal}\left(\frac{1}{\sigma}\mu - \frac{\mu}{\sigma}, \frac{1}{\sigma^2}\sigma^2\right)$$

by (**)

3 pts for solving/proving with change of variables formula

$$\sim \text{Normal}(0, 1)$$